

Is privacy compatible with truthfulness?

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Abstract

In the area of privacy-preserving data mining, a differentially private mechanism intuitively encourages people to share their data truthfully because they are at little risk of revealing their own information. However, we argue that this interpretation is incomplete because external incentives are necessary for people to participate in databases, and so data release mechanisms should not only be differentially private but also compatible with those incentives, otherwise the data collected may be false. We apply the notion of *truthfulness* from game theory. In certain settings, it turns out that existing differentially private mechanisms do not encourage participants to report their information truthfully.

On the positive side, we exhibit a transformation that takes truthful mechanisms and transforms them into differentially private mechanisms that remain truthful. Our transformation applies to games where the type space is small and the goal is to optimize an insensitive quantity such as social welfare. Our transformation incurs only a small additive loss in optimality, and it is computationally efficient. Combined with the VCG mechanism, our transformation implies that there exists a differentially private, truthful, and approximately efficient mechanism for any social welfare game with small type space.

We also study a model where an explicit numerical cost is assigned to the information leaked by a mechanism. We show that in this case, even differential privacy may not be strong enough of a notion to motivate people to participate truthfully. We show that mechanisms that release a perturbed histogram of the database may reveal too much information. We also show that, in general, any mechanism that outputs a synopsis that resembles the original database (such as the mechanism of Blum et al. (STOC '08)) may reveal too much information.

Of independent interest, one corollary of our techniques is a new lower bound on the sample complexity of differentially private non-interactive synopsis generators.

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1 Introduction

As our world becomes ever more digitized and connected, concerns about the privacy of our personal information have become increasingly pressing. With various organizations collecting, accessing, and storing individuals' data, our desire to fully take advantage of the data has come into stark tension with the equally important desire to keep information about various aspects of our lives private, including for example our medical histories and our demographic information.

Understanding and resolving this tension has been a long-standing goal in the statistics and data analysis literature [6, 16]. More recently, the theoretical computer science community has provided definitions and a rigorous treatment of this question [7, 8, 11, 4, 12], culminating in the definition and study of differential privacy [11, 8]. See [9] for a more complete overview of the area and further references.

Differential privacy guarantees the following: a (randomized) data release mechanism is ϵ -differentially private if, for any input database, any participant i in the database, and any possible output of the release mechanism s , the presence or absence of participant i causes at most a multiplicative e^ϵ change in the probability of the mechanism outputting s . The question is whether one can achieve privacy while at the same time preserving some notion of utility, since one could always build a trivially private (but not very useful) mechanism that outputs a constant value without looking at the input data. The advantage of working with the rigorous definition provided by differential privacy is that one can now show that certain tasks can be done with formal guarantees of utility and privacy [2, 4, 19, 21, 15, 23].

For a problem where one can achieve both meaningful utility and differential privacy, the differential privacy guarantee is interpreted as follows: since the output distribution of the mechanism is barely affected by the presence or absence of participant i in the above very precise and very strong sense, therefore the release mechanism leaks very little private information about participant i , and participant i should feel comfortable entering his data in the database.

Interpreting differential privacy. In this paper we investigate this interpretation. The first question we pose is why do individuals want to participate in a database at all? If individuals are indeed concerned about the privacy of their information, then they must receive some incentive to participate in order to counteract the preference for guarding their information private.

It seems hard to imagine individuals submitting private data to an organization that gives them no incentive whatsoever (interpreting “incentive” broadly). For example, individuals do not willingly submit their information to marketing companies selling something they have no interest in (unless they are incentivized by money or the chance of winning a prize). On the other hand, when something can be gained, such as the pleasure or connection to friends offered by Facebook or by using email hosted on Google, individuals do readily reveal private information. Approaching incentives from the point of view of game theory, we therefore posit that for any database where the information being collected is sensitive, there must be an associated game whose outcome incentivizes the individuals to participate in the database.

These incentives may come in various forms. Sometimes, the incentives of the individual are aligned with those of the database owner; for instance, in a medical database or a census, both the owner and the individuals share a common goal of furthering medical research or advancing public policy (in the case of the census, individuals may also want to avoid penalties). However, in general, the incentives of the database owner and individuals might diverge; for instance, in elections, the database owner (ideally) wants the candidate with the most votes to win, while individuals want their own candidate to win. In this paper we study databases where individuals' incentives may diverge from those of the database owner. This is the general case typically considered in game

theory.

Databases can be further divided into two kinds: verified and unverified. In a *verified* database, individuals cannot misreport data, for example medical databases where the information is measured by doctors and not by the individuals themselves. See *e.g.* the discussion about Ghosh and Roth [17] in the related work section for some remarks about game-theoretic considerations for verified databases. In this paper, we focus on *unverified* databases: the data is reported by individuals themselves, who may misreport their information in order to increase their utility. Again, this is the general case typically studied in game theory.

Our model: overview. In our model, a database consists of one row of information per individual. From this database an output is computed. This output includes a game-theoretic “outcome”, which determines the utility for each individual, as well as determining the “global utility” to be optimized (for example, this may be the sum of the individuals’ utilities, or it may be the revenue to the database owner). The output may also contain other auxiliary information about the database besides the outcome, for instance some kind of synopsis of the individuals’ information.

When the individuals’ private information is sensitive, we would like the output to preserve their privacy, and so we would like to build a mechanism whose output (both the outcome and any auxiliary information) is differentially private. In addition, because there is explicit utility associated with the outcome, we also require that the individuals should be incentivized to report their true data, and that the outcome should be approximately optimal with respect to the global utility. These correspond to the notions of truthfulness and efficiency, notions from mechanism design that we now discuss.

Mechanism design. Let n denote the number of individuals, also called players. Each player has a private type, which is just the game-theoretic terminology for the player’s private information. There is a global utility function that maps the players’ types plus an outcome of the game to a real number, which we will normalize to take value in $[-n, n]$. The players’ private information also gives them a *private* utility which depends on their type and the game’s outcome, which we normalize to take value in $[-1, 1]$. Each player’s goal is to optimize their private utility, and this may conflict with optimizing the global utility.

A mechanism is a (possibly randomized) function that takes a set of declared players’ types and generates an output of the game and possibly payments to be made to or received from the players. One goal of mechanism design is to come up with a mechanism that satisfies the following two conditions:

Truthfulness For all settings of the players’ true types, no player gains by misreporting their type.

Efficiency For all settings of the players’ true types, the expected global utility of the output of the mechanism should be close to the maximum possible global utility up to some additive error $\delta = o(n)$.

Clearly efficiency is an essential property of the mechanism, since one goal of the mechanism is to optimize the global utility function. (Throughout this paper, “efficiency” refers to the game-theoretic notion. Computational efficiency will be explicitly written as such.)

When the game is associated with collecting data for an unverified database, truthfulness is also an essential property, since otherwise the data being collected may be false and would be of little use to the database owner. Moreover, one could argue that privacy does not make sense without truthfulness, since protecting the privacy of false data seems superfluous.

In the game theoretic setting, truthfulness and participation are closely related: one can always extend a game to have a “ \perp ” type that represents non-participation, and have a player deviate

to the “ \perp ” type to represent non-participation. Conversely, a player who reports a type that is independent of his true type is in some sense choosing not to participate. In the rest of this paper, we will focus on truthfulness, with the understanding that non-participation is a specific kind of deviation from truthfulness.

Roadmap. Since truthfulness is essential in unverified databases, our goal in this paper is to understand the interactions that arise when differential privacy and truthfulness are studied together. We will see that mechanisms that seem satisfactory when only privacy or only truthfulness are studied in isolation are not necessarily satisfactory when the two are considered simultaneously.

We will look at the relationship between differential privacy and truthfulness in two frameworks. The first is to construct mechanisms that are (strictly) truthful and δ -efficient, while simultaneously satisfying ϵ -differential privacy (for non-trivial values of δ and ϵ). In this model, privacy is not given an explicit numerical value, but is simply an additional property of the mechanism. We call such mechanisms PTE, and we show the first PTE mechanisms by exhibiting a transformation from truthful and efficient mechanisms into PTE mechanisms for games where the type space is small and where the global utility is insensitive to any particular player’s type and depends only on the number of players who have each type.

The second framework is where the value of privacy is explicitly quantified. We will assign a non-zero cost to the amount of information leaked by a mechanism, and we ask when does the cost of revealing information overwhelm the utility a player gets from the game. We will show that when privacy has a non-zero cost, even differentially private mechanisms may not motivate individuals to reveal their true information.

While the first framework may seem overly simple by not quantifying privacy loss, the techniques developed there have subsequently been shown to apply to the second framework as well (see [Section 1.3](#)).

Previous work. There has already been a fruitful interaction between game theory and the study of differential privacy. The most closely related work is the “Exponential Mechanism” of McSherry and Talwar [21], which achieves privacy, *approximate* truthfulness, and efficiency. Approximate truthfulness says that the amount any player can gain by announcing a false type is small (but possibly positive). Approximate truthfulness is in fact a consequence of differential privacy, which says that any one player’s type can only affect the output distribution of the mechanism by a little. McSherry and Talwar [21] interpret approximate truthfulness to mean that the player might as well announce his type honestly, since he has little incentive to deviate. They also show that their mechanism achieves good efficiency (assuming the players behave truthfully).

Unfortunately, the approach suffers from the following two drawbacks, first observed by Nissim et al. [24]. First, in the Exponential Mechanism, *all* possible strategies of any individual player lead to approximately the same output distribution of the mechanism. If one accepts the interpretation that players are indifferent to small changes in their utility, then intuitively privacy may even motivate players to *lie*: players are indifferent to the small amount of utility from the game outcome that is lost by lying, but they would prefer to lie because that would reduce the amount of information leaked by the mechanism about their private type. (We will formalize this concern in [Section 5](#), see also the discussion in [Section 1.2](#).) Thus, it seems that approximate truthfulness cannot substitute for standard truthfulness when privacy is desired.

Second, Nissim et al. [24] showed that the Exponential Mechanism is *not* truthful for a game based on digital goods auctions, and therefore the relaxation to approximate truthfulness is inherently necessary. (On the other hand, the counter-example of Nissim et al. [24] is, in some sense, not

surprising because it has *no* truthful and efficient mechanism. We give a different counter-example in [Theorem 1.1](#), where the Exponential Mechanism is not truthful, but where there exists a private, truthful, and efficient mechanism.)

Nissim et al. [24] go on to show that by combining the Exponential Mechanism with a “Gap Mechanism” that incentivizes honesty, one can get a fully truthful mechanism. They apply this mechanism to give a truthful and approximately efficient mechanism for the k -facility problem on a line, which is a more general version of the 1-facility problem on a line that we will study in this paper. Unfortunately, their mechanism is not differentially private, because the Gap Mechanism relies on constraining the post-actions of the players, and this constraint reveals the types of the players in a very non-private way.

Ghosh and Roth [17] consider a question that is related but orthogonal to our work: they consider *verified* databases where each player has private information that a database owner wants to gather. Their mechanism allows each player i to declare a pricing function $c_i : \mathbb{R} \rightarrow \mathbb{R}$ such that $c_i(\varepsilon)$ represents the minimal payment that player i would require to submit his information to a database that is then published via an ε -differentially private sanitization mechanism. Their mechanism uses these pricing functions to compute a value of ε and payment values that are paid out to each player such that enough of them are incentivized to participate to make the outcome of the sanitizer accurate. In their model, the players cannot lie about their private information, only about how much they value their privacy. Our model focuses on deviations resulting from players reporting false private information, but does not explicitly consider players’ lying about their valuation of their privacy.

Feigenbaum et al. [14] study how to keep information private even from the database owner, *i.e.* before running sanitization. We do not study this problem here and treat the database owner as a trusted party. We note however that, using standard cryptographic assumptions and protocols, one can replace a trusted database owner by a secure multiparty computation among the individuals.

1.1 Private, truthful, and efficient (PTE) mechanisms

The Exponential Mechanism is not necessarily truthful. Our first result revisits the question of whether or not the Exponential Mechanism is truthful. As mentioned above, Nissim et al. [24] already exhibited a counter-example showing that the Exponential Mechanism is not truthful. However, the counter-example they give is somewhat artificial because it is easy to see that *no* efficient and truthful mechanism exists for their game (even without considering privacy).

We believe the following [Theorem 1.1](#) highlights the drawback of the Exponential Mechanism in sharper relief than the counter-example of [24], because the LINE-1-FAC game we consider *does* have a truthful and efficient mechanism. In fact, we will see that there even exists a PTE mechanism for LINE-1-FAC.

We consider the well-studied 1-facility location on a line game, denoted LINE-1-FAC, which is a special case of the k -facility problem [24, 20, 1] and also of the single-peaked preference games [22, 25]. In the LINE-1-FAC game, each player has a point on the interval $[0, 1]$ and the mechanism is supposed to output $s \in [0, 1]$. Each individual wants to minimize their distance to s , while the mechanism wants to minimize the sum of all the individuals’ distances to s . This game *does* have a truthful and efficient mechanism, namely outputting the left-most median player’s point [22]. We prove:

Theorem 1.1. *For LINE-1-FAC, the Exponential Mechanism with any privacy $\varepsilon > 0$ is not truthful.*

Transforming a truthful and efficient mechanism into a private, truthful, and efficient mechanism. We give a transformation for a large class of games that converts truthful and

efficient mechanisms into PTE mechanisms, *i.e.* truthful and δ -efficient mechanisms that also satisfy ε -differential privacy, for non-trivial ε, δ . To the best of our knowledge, these are the first PTE mechanisms exhibited for non-trivial games. Furthermore, our transformation is computationally efficient, and it applies to all games with small type space. Therefore, applying our transformation to the VCG mechanism gives a PTE mechanism for *all* social welfare games with small type space.

We note here that our definition of privacy is the relaxed version of [10], which, in addition to the e^ε multiplicative difference, also allows an additive error η , usually taken to be negligible, in the difference between the output distributions of two databases.

The transformation. The idea of the transformation is based on the ideas for privately releasing histogram data [11]. Suppose the type space of the game is a finite set of size q . For player inputs $\underline{t} = (t_1, \dots, t_n)$ where t_i is the type of the i 'th player, we will let \underline{h} denote the histogram with q entries, one for each possible type, and where $h_j = |\{i \mid t_i = j\}|$, the number of players who have type j .

Suppose each individual player's utility function is $v(t, s)$ where t is the player's type and s is an outcome of the game; suppose that v lies in the interval $[-1, 1]$. Suppose that the global utility function $w(\underline{t}, s)$ is anonymous (*i.e.* it depends only on the histogram \underline{h} and not on player identities), and is insensitive: namely if $\underline{h}, \underline{h}'$ are close in ℓ_1 distance then for all outcomes s it holds that $|w(\underline{h}, s) - w(\underline{h}', s)|$ is small. For example, the social welfare function, which is just the sum of the individual players' utilities, is an example of such a global utility function. Without loss of generality, we can assume that all mechanisms for games with such global utility functions need only depend on the histogram of player types, rather than looking at the individual players' types. (We show this in [Appendix C](#).)

At a high level, our transformation from truthful and efficient to PTE works by constructing the histogram of inputs, adding independent noise distributed according to the two-sided geometric distribution to each of the entries of the histogram, and then running the original truthful and efficient mechanism on the perturbed histogram. Care must be taken that the noise does not create negative entries into the histogram; simply truncating negative entries to 0 does not give truthful mechanisms. We show a way to achieve this in [Section 3](#). Our procedure gives the following theorem.

Theorem 1.2 (From truthful and efficient mechanisms to PTE mechanisms, informal. See [Theorem 3.3](#).)
Let $\varepsilon, \eta > 0$. Let G be a game where the type space is of size q and where the global utility function depends only on the histogram and is insensitive to individual players' types. Let M be a truthful and δ -approximately efficient mechanism. Then there exists a truthful mechanism M' for the game G that is (ε, η) -differentially private and is $(\delta + O(q \log(q/\eta)/\varepsilon))$ -approximately efficient. Furthermore, if M is computationally efficient then so is M' , and if M is moneyless then so is M' .

Application to LINE-1-FAC. We will show that using a simple rounding procedure and then applying [Theorem 1.2](#) to a mechanism for the discretized version of the LINE-1-FAC game, we can give a PTE mechanism for LINE-1-FAC.

Corollary 1.3. *For all $\varepsilon, \eta > 0$, there is a (ε, η) -differentially private, truthful, and $O(n^{1/2} \log(n/\eta)/\varepsilon)$ approximately efficient mechanism for LINE-1-FAC.*

PTE mechanism for *all* social welfare games with small type space. In fact, because the VCG mechanism gives a general truthful and perfectly efficient mechanism (using money) for all social welfare games, *i.e.* games optimizing the sum of the individuals' utilities, our main theorem implies the following:

Corollary 1.4. Fix $\varepsilon, \eta > 0$. For any game G with n players and where the type space has size q , and where the goal is to optimize social welfare, there is a truthful mechanism for G that is (ε, η) -differentially private and is $O(q \log(q/\eta)/\varepsilon)$ -approximately efficient.

1.2 The value of privacy

Prior work (including the results outlined in the previous section) studied privacy and utility as orthogonal features. One posited a utility measure and showed that one could achieve a truthful and efficient solution that simultaneously satisfied differential privacy. However, if one really believes that participants value their privacy, then this should explicitly be taken into account in their utility functions. Not only does this allow us to quantify *how much* participants value their privacy, it also allows us to model tradeoffs between utility and privacy. That is, participants may be willing to sacrifice some of the utility they reap from a game by reporting a false type and thereby protecting their privacy. Vice versa, participants may be willing to sacrifice their privacy if by doing so they gain a larger value from the outcome of the game.

We show that it is possible even for PTE mechanisms to leak too much information to achieve truthfulness in games where there exists a tradeoff between utility and privacy.

Quantifying privacy The first task is to define a measure of how much information a mechanism leaks. One natural criterion is that the information cost should be non-negative. Another natural criterion is that the information cost should be 0 if a player reports a type that is independent of his honest type. This criterion implies that information cost cannot *solely* be a function of the player’s type and the outcome of the game, because the notion “independent of the honest type” is inherently a statement about how the player behaves on *all* possible values of his type. As a thought experiment, fix any $t \in Q$, and consider the following two strategies: first is the truthful strategy, and second the strategy that always outputs t regardless of what the player’s actual type is. Then, in the case that the player’s actual type is t , the two strategies give exactly the same output, but intuitively the truthful strategy might reveal information about the player’s type while the constant t strategy does not.

Therefore our measure of information cost cannot be expressed as a modification of a “traditional” utility function that depends solely on the player’s type and outcome. Instead, we work with a measure that depends on the player’s strategy over all possible types. Since we do not know in general what features of the private type are important to protect (this depends on the context and application), we use the following general information-theoretic measure, which is essentially a “max-divergence” measure between the output distributions of the mechanism on different player inputs (shifted slightly to be non-negative). The max-divergence was introduced in the context of differential privacy by Dwork et al. [13], and we refer the reader there for a discussion. We work with the η -approximate notion here in order to prevent the information cost from being artificially large (or even infinite) due to some low-probability events.

Definition 1.5. The η -approximate information cost of strategy σ to player i on input \underline{t} and for mechanism M is

$$\text{IC}_M^\eta(\sigma, \underline{t}, i) = \max_{t' \in Q} \max_{\substack{B \subseteq S \times \mathbb{R}^n \\ \Pr[M(\underline{t}^{-i}, \sigma(t)) \in B] > \eta}} \log \frac{\Pr[M(\underline{t}^{-i}, \sigma(t)) \in B] - \eta}{\Pr[M(\underline{t}^{-i}, \sigma(t')) \in B](1 - \eta)}$$

We write IC_M to denote IC_M^0 .

σ is a “strategy”, a randomized function mapping types to types; the identity function Id is the truthful strategy. If $\sigma(t)$ does not depend on t , then $\text{IC}_M^\eta(\sigma, \underline{t}, i) = 0$ always. A simple calculation shows that if a mechanism is (ε, η) -differentially private, then it holds that $\text{IC}_M^\eta(\sigma, \underline{t}, i) \leq \varepsilon + \log \frac{1}{1-\eta}$ for all σ, \underline{t}, i . Note that $\log \frac{1}{1-\eta} \approx 0$ since we will take η to be negligible.

Tradeoffs between the value of the game and the value of privacy. Given the definition of information cost defined above, we would like to explicitly incorporate it into the utility of the player. Let us use the word “game value” or simply “value” and the letter v to denote the benefit that the player extracts from the outcome of the game without taking into account privacy, and we will let overall utility, denoted by u , denote the expected game value minus the information cost, weighted with a factor ν_i that expresses how much the individual values his privacy relative to the value of the game:

$$u_i^\eta(\sigma, \underline{t}) = \mathbb{E}_{M, \sigma}[v(t_i, M(\sigma(\underline{t}^{-i}, t_i)))] - \nu_i \text{IC}_M^\eta(\sigma, \underline{t}, i) \quad (1.1)$$

We write u_i to denote u_i^0 .

1.2.1 Releasing histograms for the LINE-1-FAC game.

Our first negative result draws again on the LINE-1-FAC game. [Corollary 1.3](#) says that a PTE mechanism for LINE-1-FAC exists, let us denote it by M (see [Algorithm 4.6](#) for the definition of the mechanism). Roughly, M functions by rounding the positions of the players to discrete points, constructing a histogram of the players’ rounded locations, perturbing the histogram, and finally outputting the median point of the perturbed histogram. Suppose now that, instead of outputting the median, the mechanism outputs “more than necessary” and also publishes the entire perturbed histogram of the players’ locations as well. Call this mechanism \hat{M} . We believe that \hat{M} , represents a plausible situation in the real world: an agency gathers data for a single explicit purpose, for example to decide where to build a hospital so as best to serve local residents, but then publishes the entire perturbed histogram data so that it may be of use to other agencies that may want to use it in a different way.

\hat{M} is actually PTE: truthfulness and efficiency remain because the facility location output is the same as M , while privacy also remains because the perturbation of the histogram was designed to render the entire histogram differentially private. However, we show in [Theorem 5.5](#) that if we set the parameters so that the mechanism is $(2\varepsilon, \eta)$ -differentially private, then the information cost is greater than ε . On the other hand, we show in [Theorem 5.2](#) that there exist situations where the amount of value that a player loses by ignoring his true type and declaring a fixed type, say 0, is at most $e^{-\Omega(n)}$. This implies the following:

Theorem 1.6 (Informal, see [Corollary 5.7](#)). *The PTE mechanism for LINE-1-FAC given by \hat{M} is untruthful when one takes into account the information cost, as long as some player i gives weight $\nu_i > e^{-\Omega(n)}$ to the information cost.*

The hypothesis that $\nu_i > e^{-\Omega(n)}$ is essentially always true, see [Remark 5.1](#) for a discussion. The above theorem therefore says that one should not publish the entire histogram, otherwise players may be incentivized to lie.

1.2.2 Releasing synopses.

We prove a more general theorem about the information cost of any mechanism that publishes information that can be useful for many different count queries: if Q is the type space of the players and $F : Q \rightarrow \{0, 1\}$, then let $\bar{F}(\underline{t}) = \frac{1}{n} \sum_{i=1}^n F(t_i)$.

Definition 1.7. M is a (γ, ρ) -synopsis generator on n -player inputs with respect to a class \mathcal{C} if there is a real-valued function $P(s, F)$ such that, for all $\underline{t} \in Q^n$, $\Pr[\max_{F \in \mathcal{C}} |\overline{F}(\underline{t}) - P(s, F)| \leq \gamma] \geq 1 - \rho$.

Blum et al. [4] showed that, if $n = \tilde{O}(\frac{d \log |Q|}{\varepsilon \gamma^3})$ where d denotes the VC-dimension of \mathcal{C} , then there exists M that is a (γ, ρ) -synopsis generator with respect to \mathcal{C} and also ε -differentially private. We show a lower bound on the amount of information such a mechanism must leak.

Theorem 1.8. Fix any $\gamma \in (0, \frac{1}{5}), \rho \in (0, 1)$. Suppose M is a mechanism that is a (γ, ρ) -synopsis generator on n -player inputs with respect to \mathcal{C} , which has VC dimension d . Then for all $\underline{t} \in Q^n$, there exists $\underline{t}' \in Q^n$ and $i \in [n]$ such that $\underline{t}, \underline{t}'$ differ in at most $4\gamma n$ entries and such that $\text{IC}_M(\text{Id}, \underline{t}', i) \geq \min(\Omega(\frac{d}{\gamma n}), \Omega(1))$.

This shows that synopsis generators such as the one proposed by Blum et al. [4] must inherently reveal a lot of information some player’s type. Furthermore, it must leak information on “many” databases: for any database, there is a nearby one with a player with high information cost. Interestingly, the proof uses a construction of a combinatorial design to show that, for every \underline{t} , there exists an exponentially large family of \underline{t}' that differ little from \underline{t} such that one of them has a player with large information cost.

This theorem means the following: consider a mechanism M that releases a synopsis good for \mathcal{C} with VC-dimension d . M must also compute an outcome of some associated game in order to incentivize participation. If the mechanism M has the property that an individual’s value diminishes by very little if he mis-reports his type, then Theorem 1.8 says that the individual will prefer to lie and declare, say, a constant value because his gain in information cost will compensate for his loss in value. This holds even if M is differentially private. As a specific instantiation, we prove that if a mechanism M approximately solves the 1-facility location problem over any metric space and simultaneously releases a synopsis, then M cannot be truthful. See Section 5.2 for formal statements.

Lower bound on sample complexity. For any ε -differentially private mechanism, the information cost is upper bounded by ε . Therefore Theorem 1.8 implies the following corollary. (Note that this is incomparable to the previously known lower bound of [3], who showed that $n \leq d/2$ implies $\varepsilon \geq \Omega(\frac{1}{\gamma})$.)

Corollary 1.9. Fix any $\gamma \in (0, \frac{1}{5}), \rho \in (0, 1)$. If M is an ε -differentially private (γ, ρ) -synopsis generator for a class \mathcal{C} of VC-dimension d for a database of size n , then $\varepsilon \geq \min(\Omega(\frac{d}{\gamma n}), \Omega(1))$.

1.3 Comments and subsequent work

Subsequent to the initial version of this manuscript, Chen et al. [5] showed that it is possible to build truthful mechanisms even when the players’ utilities explicitly take into account the cost of information leaked (for example, using the measure of Equation 1.1), thus answering affirmatively one of the open questions posed in the initial version of this manuscript. Their result holds for a general class of information costs, including but not limited to Definition 1.5. One of the ingredients in their work is the TE-to-PTE transformation (Theorem 3.3) presented in this paper.

2 Preliminaries

We let $[q] = \{1, \dots, q\}$. We identify sets S with the uniform distribution over the set. For a distribution X , we write $x \leftarrow_{\mathbb{R}} X$ to denote a random sample from that distribution. For $x \in \mathbb{R}^n$,

we let $\|\underline{x}\|_1 = \sum_{i=1}^n |x_i|$ denote the ℓ_1 norm and $\|\underline{x}\|_\infty = \max_i |x_i|$ the ℓ_∞ norm. Additional definitions and facts appear in the appendix.

Mechanisms. A game with n players consists of the type space Q , a space S of possible outcomes of the game, a valuation function $v : Q \times S \rightarrow [-1, 1]$, and a global utility function $w : Q^n \times S \rightarrow \mathbb{R}$. When Q is finite, we let $|Q| = q$. For $i \in [n]$, for t_1, \dots, t_n , we let \underline{t}^{-i} denote the vector with $n-1$ entries given by $t_1, \dots, t_{i-1}, t_{i+1}, \dots, t_n$. Define $\underline{h}(\underline{t}) \in \mathbb{Z}^q$ the histogram of the input, where $h_j = |\{i \mid t_i = j\}|$. We assume *w.l.o.g.* that the type space has a special \perp type, and any player that has this special type is ignored. (Namely, an input with n players, k of whom have type \perp , is treated as an input to the game with $n-k$ players with those k players removed.)

A game has an *anonymous* global utility function w if there is a function w' such that for all \underline{t}, s , it holds that $w(\underline{t}, s) = w'(\underline{h}(\underline{t}), s)$. We abuse notation and write $w(\underline{h}, s)$ when we mean $w(\underline{t}, s)$ where $\underline{h} = \underline{h}(\underline{t})$. An anonymous global utility function is *insensitive* if for all integers k , $\|\underline{h} - \underline{h}'\|_1 \leq k$ implies that for all s , $|w(\underline{h}, s) - w(\underline{h}', s)| \leq k$. The *social welfare* is $w(\underline{t}, s) = \sum_{i=1}^n v(t_i, s)$. It is anonymous and insensitive.

A mechanism M is a (randomized) function that takes types t_1, \dots, t_n for all of the players and samples an outcome $(s, p_1, \dots, p_n) \leftarrow_{\mathbb{R}} M(t_1, \dots, t_n)$ where $s \in S$ is the outcome and $p_i \in \mathbb{R}$ are the payoffs. We say that M is *moneyless* if $p_1 = \dots = p_n = 0$ for all inputs and random coins. We are concerned with asymptotic analysis, so a well-defined mechanism M must be able to handle any number of players n .

We say a mechanism M is *truthful* if for each player $i \in [n]$, and for all $\underline{t}^{-i}, t_i, t'_i$, it holds that $\mathbb{E}_{(s,p) \leftarrow_{\mathbb{R}} M(\underline{t})} [v(t_i, s) - p_i] \geq \mathbb{E}_{(s,p) \leftarrow_{\mathbb{R}} M(\underline{t}^{-i}, t'_i)} [v(t_i, s) - p_i]$. This must be slightly generalized to handle the tradeoff utility of Equation 1.1, which depends on the player's strategy σ : if a player uses strategy σ , on true type t he samples from $\sigma(t)$ and reports the sample as his type to the mechanism. For this notion of utility, we say that M is truthful if for all $i \in [n]$, all strategies σ and all inputs \underline{t} , it holds that $u_i^\eta(\text{Id}, \underline{t}) \geq u_i^\eta(\sigma, \underline{t})$.

We say a mechanism is $\delta(n)$ -*efficient* if for all inputs \underline{t} on n players, it holds that $\mathbb{E}_M [w(\underline{t}, M(\underline{t}))] \geq \max_{s \in S} w(\underline{t}, s) - \delta(n)$. Observe that our definition of efficiency allows for an *additive* error. In contrast, most work in the mechanism design literature on approximate mechanisms deal with *multiplicative* error. However, additive error is more suitable when working with differential privacy.

Differential privacy A mechanism M is (ε, η) -*differentially private* if for all $i \in [n]$, all $\underline{t} \in Q^n$, all $\underline{t}' \in Q$, and all $U \subseteq S \times \mathbb{R}^n$, it holds that $\Pr[M(\underline{t}) \in U] \leq e^\varepsilon \Pr[M(\underline{t}^{-i}, \underline{t}') \in U] + \eta$. Typically we think of $\varepsilon > 0$ being a small constant and η as being $o(1)$, preferably negligible. We will frequently let $\alpha = e^{-\varepsilon}$.

Let \mathcal{G}_ε denote the geometric distribution, with probability mass function $f(x) = \frac{1-e^{-\varepsilon}}{1+e^{-\varepsilon}} e^{-\varepsilon|x|}$. We will also use the following distribution:

Definition 2.1. Distribution $\mathcal{H}_{\varepsilon, \tau, q}$: sample $\underline{\zeta}' \leftarrow_{\mathbb{R}} \mathcal{G}_\varepsilon^q$. If $\|\underline{\zeta}'\|_\infty > \tau$, output 0, otherwise output $\underline{\zeta}$.

We prove the following lemma in the appendix.

Lemma 2.2. For all $i, j \in [q]$ and $U \subseteq \mathbb{Z}^q$, for $\underline{\zeta}'$ sampled from $\mathcal{H}_{\varepsilon, \tau, q}$ it holds that:

$$\Pr[\underline{\zeta}' \in U] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' \in U] + \frac{2q\alpha^\tau}{1+\alpha}$$

where e_i denotes the i 'th standard basis vector.

It is straightforward to calculate that for $\zeta \leftarrow_{\mathbb{R}} \mathcal{G}_\varepsilon$:

$$\Pr[|\zeta| \geq \tau] = \frac{2\alpha^\tau}{1+\alpha} \tag{2.1}$$

Negative binomial distribution. The negative binomial distribution $\mathcal{NB}_{q,\varepsilon}$ is defined using the probability mass function

$$f(x) = \binom{x+q-1}{q-1} \cdot (1-e^{-\varepsilon})^q e^{-\varepsilon x}$$

We also note the following facts:

Fact 2.3 (e.g. [26]). 1. The sum of q two-sided random geometric variables $\sum_{j=1}^q \zeta_j$, where each ζ_j is distributed according to \mathcal{G}_ε , is distributed identically to $Y - Y'$ where both Y, Y' are independent and identically distributed according to the $\mathcal{NB}_{q,\varepsilon}$.¹

2. Suppose Y is distributed according to $\mathcal{NB}_{q,\varepsilon}$. Then $\Pr[Y \geq t] = \Pr[Z \leq q]$ where Z is a binomial random variable for an experiment with $q+t$ trials and success probability $(1-e^{-\varepsilon})$ for each trial.

3 PTE Mechanisms

Definition 3.1. The LINE-1-FAC game is defined as follows. The player types are $t_i \in [0, 1]$. The outcome of the game is a point $s \in [0, 1]$. The utility function is $v(t, s) = -|t - s|$ and the global utility is the social welfare: $w(\underline{t}, s) = \sum_{i \in [n]} v(t_i, s) = -\sum_{i \in [n]} |t_i - s|$.

The moneyless mechanism that outputs the median (breaking ties say by picking the left median point) of the $\{t_1, \dots, t_n\}$ is truthful and achieves optimal social welfare [22].

Theorem 1.1 (Restated). For LINE-1-FAC, the Exponential Mechanism with any privacy $\varepsilon > 0$ is not truthful.

Proof. For $n = 2$, set $t_1 = 0, t_2 = 2/3$. (This can easily be extended to an arbitrary even number of players n by placing $n/2$ players at 0 and $(n/2 - 1)$ players at 1, and one player at $2/3$.)

Claim 3.2. For all $\varepsilon > 0$, if player 2 declares 1 then his utility under the Exponential Mechanism is $-5/18$, while if he declares $2/3$ then his utility is strictly less than $-5/18$.

Proof of Claim 3.2. The Exponential Mechanism M_{Exp} generates an output s according to the density function $f(s) = \frac{e^{\varepsilon w(\underline{t}, s)}}{\int_0^1 e^{\varepsilon w(\underline{t}, s)} ds}$. One can compute that the expected utility of player 2 if he reports 1 is $-\int_0^1 |2/3 - s| ds = -5/18$, since in this case there are exactly half the players at 0 and at 1 and so the social welfare function (and hence the mechanism's output) is uniform over $[0, 1]$.

If player 2 is truthful, then the social welfare equals $w(\underline{t}, s) = -s - |2/3 - s|$, and therefore the expected utility of player 2 is:

$$V \stackrel{\text{def}}{=} \mathbb{E}_{s \leftarrow R_{M_{\text{Exp}}(0, 2/3)}} [v(2/3, s)] = \frac{-\int_0^1 e^{-\varepsilon(s+|2/3-s|)} |2/3 - s| ds}{\int_0^1 e^{-\varepsilon(s+|2/3-s|)} ds} \quad (3.1)$$

We claim that for all $\varepsilon > 0$, it holds that $V < -5/18$. This is equivalent to proving

$$0 < -V - 5/18 = \frac{\int_0^1 e^{-\varepsilon(s+|2/3-s|)} \cdot (|2/3 - s| - 5/18) ds}{\int_0^1 e^{-\varepsilon(s+|2/3-s|)} ds}$$

¹This follows from the fact that a two-sided geometric random variable is identical to the difference between two one-sided geometric random variables, and the sum of one-sided geometric random variables is a negative binomial.

Observe that the denominator is positive, so it suffices to prove that the numerator is positive for all $\varepsilon > 0$. Evaluating the integral we obtain that

$$\int_0^1 e^{-\varepsilon(s+|2/3-s|)} \cdot (|2/3-s| - 5/18) ds = e^{-2\varepsilon/3} \left(\frac{1}{27} + \frac{9 - 5\varepsilon - e^{-2\varepsilon/3}(\varepsilon + 9)}{36\varepsilon^2} \right) \quad (3.2)$$

Since $e^{-2\varepsilon/3} > 0$ so we can multiply the LHS of Equation 3.2 by $36\varepsilon^2 e^{2\varepsilon/3}$ and simplify, and it suffices to prove that

$$\frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - e^{-2\varepsilon/3}(\varepsilon + 9) > 0 \quad (3.3)$$

For large $\varepsilon \gg 0$ the RHS of Equation 3.3 is dominated by the quadratic term. For instance it is easy to check that for all $\varepsilon > 3$, it holds that $e^{-2\varepsilon/3} < 1/7$ and so:

$$\frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - e^{-2\varepsilon/3}(\varepsilon + 9) > \frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - \frac{\varepsilon}{7} - 9/7 \quad (3.4)$$

$$> \frac{4\varepsilon^2}{3} - 5.2\varepsilon + 7.7 \quad (3.5)$$

$$> 0 \quad (3.6)$$

where the final inequality can be deduced by the fact that the quadratic polynomial in Equation 3.5 has no real roots and is non-negative. Therefore we can restrict our attention to the case $\varepsilon \leq 3$. Using the Taylor expansion of the exponential, we know that $e^{-2\varepsilon/3} \leq 1 - \frac{2\varepsilon}{3} + \frac{2\varepsilon^2}{9} - \frac{4\varepsilon^3}{81} + \frac{2\varepsilon^4}{243}$, therefore Equation 3.3 can be rewritten as:

$$\begin{aligned} \frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - e^{-2\varepsilon/3}(\varepsilon + 9) &\geq \frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - \left(1 - \frac{2\varepsilon}{3} + \frac{2\varepsilon^2}{9} - \frac{4\varepsilon^3}{81} + \frac{2\varepsilon^4}{243}\right)(\varepsilon + 9) \\ &= \frac{4\varepsilon^2}{3} + 9 - 5\varepsilon - \varepsilon - 9 + \frac{2\varepsilon^2}{3} + 6\varepsilon - \frac{2\varepsilon^3}{9} - 2\varepsilon^2 + \frac{4\varepsilon^4}{81} + \frac{4\varepsilon^3}{9} - \frac{2\varepsilon^5}{243} - \frac{2\varepsilon^4}{27} \\ &= \frac{2\varepsilon^3}{9} - \frac{2\varepsilon^4}{81} - \frac{2\varepsilon^5}{243} \\ &> \varepsilon^3 \left(\frac{2}{9} - \frac{2}{27} - \frac{2}{27} \right) \\ &> 0 \end{aligned}$$

where in the last two lines we use the fact that $\varepsilon \leq 3$. ■

3.1 PTE achievable

We now exhibit a generic transformation that converts a truthful and efficient mechanism M into a PTE mechanism M' . We assume M satisfies the following: on input \underline{t} , M computes its output depending only on $\underline{h} = \underline{h}(\underline{t})$ (and never looks at the entries of \underline{t} individually). (This is without loss of generality for anonymous games, see Appendix C.)

Theorem 3.3. *Let G be a game with type space of size q and whose global utility function is anonymous and insensitive. Suppose G has a truthful and δ -efficient histogram mechanism M .*

Then for all $\varepsilon, \eta > 0$, G also has a $(2\varepsilon, \eta)$ -differentially private, truthful, and δ' -efficient mechanism M' , where $\delta'(n) = \delta(n + O(q \log(q/\eta)/\varepsilon)) + O(q \log(q/\eta)/\varepsilon)$. M' is given by Algorithm 3.4.

Observe that if M is computationally efficient then so is M' and if M is moneyless then so is M' .

Proof of Theorem 3.3. Let M be the truthful and δ -efficient mechanism. By Equation 2.1, we can set $\tau = O(\log(q/\eta)/\varepsilon)$ such that

$$\Pr_{\zeta \leftarrow \mathcal{R}G_\varepsilon^q} [\|\underline{\zeta}\|_\infty > \tau] = \eta \quad (3.7)$$

In Algorithm 3.4 we construct a mechanism M' (using τ set as just mentioned) that is private, truthful, and δ' -approximately efficient. We prove that M' satisfies all three properties:

Truthfulness. Fix an input \underline{t} . We claim that for every choice of $\underline{\zeta}$, the mechanism is truthful. We run M on the histogram $\underline{h}' = \underline{h} + \underline{\zeta} + \tau \cdot \underline{1}$. Because $\underline{\zeta}$ is sampled from $\mathcal{H}_{\varepsilon, \tau, q}$, it holds for every $t \in Q$ that $h'_t \geq h_t \geq 0$. Suppose that, for this fixing of $\underline{\zeta}$, there is a deviation that benefits some player, *i.e.* there exists i and t' such that

$$\mathbb{E}[v(t_i, M(\underline{h}'))] < \mathbb{E}[v(t_i, M(\underline{h}' - e_{t_i} + e_{t'}))]$$

Then this is also a deviation on the input \underline{h}' for the original M , which contradicts the fact that M is truthful.

Efficiency. Let $\underline{\zeta}$ be sampled from $\mathcal{H}_{\varepsilon, \tau, q}$ and observe that $\|\underline{\zeta}\|_{\infty} \leq \tau$. For any input \underline{t} define the modified histogram $\underline{h}' = \underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1}$. Observe that $|\underline{h} - \underline{h}'|_1 \leq 2q\tau$. By the insensitivity of w , it holds that for all possible outcomes of the game s' , it holds that $|w(\underline{h}', s') - w(\underline{t}, s')| \leq 2q\tau$. Therefore it holds that:

$$\mathbb{E}_M[w(\underline{h}, M(\underline{h}'))] \geq \mathbb{E}_M[w(\underline{h}', M(\underline{h}'))] - 2q\tau \quad (3.8)$$

$$\geq \max_{s'} w(\underline{h}', s') - 2q\tau - \delta(n + 2q\tau) \quad (3.9)$$

$$\geq w(\underline{h}', s_0) - 2q\tau - \delta(n + 2q\tau) \quad (3.10)$$

$$\geq \max_s w(\underline{h}, s) - 4q\tau - \delta(n + 2q\tau) \quad (3.11)$$

In [Equation 3.9](#) we use the efficiency of M and the fact that \underline{h}' corresponds to a game with at most $n + 2q\tau$ players. We also assumed that $\delta(n) \in [0, 2n]$ is non-decreasing, which holds without loss of generality since any δ not satisfying the condition can be altered by artificially increasing δ at some points. In [Equation 3.11](#) s_0 denotes the outcome that maximizes $w(\underline{h}, s)$.

Applying [Equation 3.11](#) to the expected utility of M' , one obtains the following:

$$\mathbb{E}_{M'}[w(\underline{t}, M'(\underline{h}))] = \mathbb{E}_{\underline{\zeta} \leftarrow \mathcal{R}\mathcal{H}_{\varepsilon, \tau, q}}[w(\underline{t}, M(\underline{h}'))] \geq \max_s w(\underline{t}, s) - \delta(n + 2q\tau) - 4q\tau$$

where $\underline{h}' = \underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1}$.

Privacy. Observe that [Algorithm 3.4](#) is just perturbing the histogram according to the distribution $\mathcal{H}_{\varepsilon, \tau, q}$. Therefore, from [Lemma 2.2](#), we know that for any adjacent $\underline{h}, \underline{h}^*$, it holds that for all subsets $U \subseteq \mathbb{Z}^q$ that

$$\Pr[\underline{h} + \underline{\zeta}' \in U] \leq e^{2\varepsilon} \Pr[\underline{h}^* + \underline{\zeta}' \in U] + \frac{2qe^{-\varepsilon\tau}}{1+e^{-\varepsilon}}$$

Plugging in our choice of τ , this shows the mechanism is $(2\varepsilon, \eta)$ -differentially private. ■

Note that one popular alternative way of perturbing the histogram, by adding noise according to $\mathcal{G}_{\varepsilon}$ to each bin of the histogram and then truncating negative bins to 0, does not seem to guarantee truthfulness. The problem is that the truncation process is non-linear, and so it is unclear how to prove that a deviation in the perturbed game implies a deviation in the original game.

Corollaries. We prove [Corollary 1.3](#) by applying [Theorem 3.3](#) to a discretized version of this truthful and efficient mechanism, see [Section 4](#).

[Corollary 1.4](#), stating that all social welfare games with small type space have a PTE mechanism, follows from applying [Theorem 3.3](#) to the VCG mechanism (which is truthful and perfectly efficient for all social welfare games).

Input: types $t_1, \dots, t_n \in [q]$. Auxiliary input: ε, η privacy parameters. Set $\tau = O(\log(q/\eta)/\varepsilon)$.

1. Sample $\underline{\zeta} \leftarrow_{\text{R}} \mathcal{H}_{\varepsilon, \tau, q}$
2. Construct $\underline{h}' = \underline{h} + \underline{\zeta} + \tau \cdot \underline{1}$, where $\underline{1}$ is the all 1 vector.
3. Output $M(\underline{h}')$.

Algorithm 3.4. PTE mechanism based on a truthful and efficient mechanism M .

Input: histogram of player types \underline{h} . Let $n = \sum_{j=1}^q h_j$.

1. Output the minimal $s \geq 1$ such that $\sum_{j=1}^s h_j \geq \frac{n}{2}$.

Algorithm 4.2. Truthful and efficient mechanism for D-L1F

4 LINE-1-FAC has a PTE mechanism

We show that although the Exponential Mechanism is not truthful for LINE-1-FAC, there does exist a mechanism that is truthful, differentially private, and approximately efficient. This mechanism is given in [Algorithm 4.6](#). The idea is to reduce LINE-1-FAC to a game D-L1F (“discrete one-facility on a line”) with a small type space, and then to give a private, truthful, and efficient mechanism for D-L1F using [Theorem 3.3](#).

4.1 The D-L1F game

Definition 4.1. The D-L1F $_{\gamma}$ game is defined as follows. Assume that $\gamma > 0$ is such that $q = 1/\gamma + 1$ is an integer. The player types are $t_i \in [q]$. Output of mechanism is $s \in [q]$. Utility function is $v(t, s) = -\gamma|t - s|$. Global utility is social welfare: $w(\underline{t}, s) = \sum_{i=1}^n v(t_i, s)$.

Theorem 4.3. *There is a truthful and perfectly efficient mechanism for D-L1F $_{\gamma}$.*

The mechanism is given in [Algorithm 4.2](#). We may apply [Theorem 3.3](#), we obtain:

Corollary 4.4. *For all $\varepsilon, \eta > 0$, D-L1F $_{\gamma}$ has a (ε, η) -differentially private, truthful, and δ -efficient mechanism for $\delta = O(\log(\frac{1}{\gamma\eta})/(\gamma\varepsilon))$.*

The proof of [Theorem 4.3](#) is a special case of the proof that the median mechanism is truthful and efficient for single-peaked preferences. For completeness, we give a proof in [Section B.1](#).

4.2 Using D-L1F $_{\gamma}$ to give a PTE mechanism for LINE-1-FAC

Theorem 4.5. *For any $\gamma, \varepsilon, \eta > 0$, the mechanism of [Algorithm 4.6](#) is $(2\varepsilon, \eta)$ -differentially private, truthful, and δ -efficient for the LINE-1-FAC game, where $\delta = n\gamma + O(\frac{1}{\gamma\varepsilon} \log(\frac{1}{\gamma\eta}))$.*

As an example setting, pick $\gamma = n^{-1/2}$, which implies $\delta = O(n^{1/2} \log(n/\eta)/\varepsilon)$.

Input: player types $t_1, \dots, t_n \in [0, 1]$. Assume for simplicity that $1/\gamma$ is an integer.

1. Discretize $[0, 1]$ into $q = 1/\gamma + 1$ intervals: $[0, \gamma/2), [\gamma/2, (1 + 1/2)\gamma), \dots, [(j - 1/2)\gamma, (j + 1/2)\gamma), \dots, [(q - 1/2)\gamma, 1]$.
2. Assign player i the proxy type $\hat{t}_i = j$ such that t_i falls into the j 'th interval.
3. Let M denote the mechanism for D-L1F $_\gamma$ given by [Theorem 4.3](#). Let M' be the corresponding $(2\varepsilon, \eta)$ -differentially private, truthful, and efficient mechanism, given by [Corollary 4.4](#). Run $M'(\hat{t}_1, \dots, \hat{t}_n)$ to obtain $s \in [q]$.
4. Output γs .

Algorithm 4.6. $(2\varepsilon, \eta)$ -differentially private, truthful, and efficient mechanism for LINE-1-FAC.

Proof of [Theorem 4.5](#). $(2\varepsilon, \eta)$ -differential privacy follows immediately from [Theorem 3.3](#). Efficiency is also straightforward, because the overall error is at most the discretization error, which is bounded by γ for each player, plus the error from M' , which is bounded by $O(\frac{1}{\gamma\varepsilon} \log(\frac{1}{\gamma\eta}))$.

Truthfulness. Truthfulness must be argued more carefully because the rounding process might cause unexpected problems. By symmetry, it suffices to consider player 1.

Fix t_1, \underline{t}^{-1} , and t^* . We will show that the player can gain no utility from declaring t^* when his actual type is t_1 . Fix any choice of random coins $\underline{\zeta}$ used by M' . Let $\underline{h}' = \underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1}$ and $\underline{h}^* = \underline{h}(\underline{t}^{-1}, t^*) + \underline{\zeta} + \tau \cdot \underline{1}$. Let $s = M(\underline{h}')$ and $s^* = M(\underline{h}^*)$. From the proof of truthfulness of [Algorithm 4.2](#), we observe that M has the following property (which is stronger than truthfulness): either $s = s^*$ or $|\hat{t}_1 - s| \leq |\hat{t}_1 - s^*| - 1$. (Namely, it is impossible for $s^* \neq s$ and yet $|\hat{t}_1 - s| = |\hat{t}_1 - s^*|$.)

In the case where $s = s^*$ then there is clearly no advantage to lying, so suppose $|\hat{t}_1 - s| \leq |\hat{t}_1 - s^*| - 1$. Observe that the rounding process guarantees that for all $t \in [0, 1]$, $|t - \gamma\hat{t}| \leq \gamma/2$. Therefore we may write

$$\begin{aligned} |t_1 - \gamma s| &\leq \gamma |\hat{t}_1 - s| + \gamma/2 \\ &\leq \gamma (|\hat{t}_1 - s^*| - 1) + \gamma/2 \\ &\leq |t_1 - \gamma s^*| \end{aligned}$$

This again implies that there is no advantage to declaring t^* , and so since for every choice of $\underline{\zeta}$ the mechanism is truthful, it follows that the overall mechanism is truthful. ■

5 The value of privacy

We now explore a model that assigns non-negative costs to the information leaked by mechanisms about the private types of the players. The definition of information cost was given in [Definition 1.5](#), and the utility expressing a tradeoff between game value and information cost in [Equation 1.1](#). The results of this section show that even a differentially private mechanism may reveal too much information about individuals and thereby motivate them to lie.

Remark 5.1. In [Equation 1.1](#), ν_i models the amount of weight player i attaches to his privacy, and we assume that the ν_i are fixed and known (Ghosh and Roth [17] study mechanisms where

these valuations are private). Intuitively $1/\nu_i$ represents how many bits must be leaked for player i to lose a constant amount of utility. $\nu_i = \Omega(1/\log |Q|)$ is a realistic setting, and would mean a constant cost is incurred if a constant *fraction* of bits is leaked. Significantly smaller values of ν_i would model a situation in which the player assigns significant cost *only* when his type is essentially *completely* revealed by the mechanism. In particular, we may safely assume that the weights to satisfy $\nu_i = 2^{-o(n)}$, since otherwise the amount of utility that the player places on privacy is so small that explicitly modelling the value of privacy loses relevance.

5.1 Releasing histograms

Let M' be the PTE mechanism for LINE-1-FAC given in [Algorithm 4.6](#) (one could also just consider the discrete version, [Algorithm 4.2](#)) run with (ε, η) -differential privacy. We show first that, on certain databases, no single player can heavily influence the outcome of M' . $(2\varepsilon, \eta)$ -differential privacy implies that being untruthful can hurt the value of a player by at most $2\varepsilon + \eta$. We show that for this particular mechanism M' , there exist inputs for which the loss is much smaller.

Theorem 5.2. *Fix any $\varepsilon > 0$ and any $\eta > 0$ such that $\eta = 2^{-o(\sqrt{n}/\log n)}$ (i.e. η is not too small). Then for n large enough, for all i , there exists $\underline{t}^{-i} \in [0, 1]^{n-1}$ such that for all $t_i, t'_i \in [0, 1]$, it holds that:*

$$\mathbb{E}_{s \leftarrow R M'(\underline{t})}[v(t_i, s)] - \mathbb{E}_{s \leftarrow R M'(\underline{t}^{-i}, t'_i)}[v(t_i, s)] \leq e^{-(1-e^{-\varepsilon})^2 n} \quad (5.1)$$

Proof. To simplify our notation, suppose that the number of players is $n+1$ rather than n . By our choice of γ , [Algorithm 4.6](#) divides $[0, 1]$ into $q = n^{1/2}$ intervals. The setting of player types \underline{t}^{-1} is simply to put n players at location 0.

Recall that M' functions by building a histogram of the players' locations, generating noise distributed according to $\mathcal{H}_{\varepsilon, \tau, q}$ where $\tau = O(\log(q/\eta)/\varepsilon)$, adding $\underline{\zeta} + \tau \cdot \underline{1}$ to the histogram, and then running the deterministic mechanism M of [Algorithm 4.2](#). Notice that for our choice of \underline{t}^{-1} , there is no rounding necessary. Let us rewrite the LHS of [Equation 5.1](#) as:

$$\mathbb{E}_{\underline{\zeta}}[v(t_1, M(\underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1})) - v(t_1, M(\underline{h}(\underline{t}^{-1}, t'_1) + \underline{\zeta} + \tau \cdot \underline{1}))] \quad (5.2)$$

The main observation is that for most choices of $\underline{\zeta}$, M gives the same output regardless of what player 1 declares as its value. To state this more formally, define

$$\underline{h}^{-1} = \underline{h}(\underline{t}^{-1}) + \underline{\zeta} + \tau \cdot \underline{1} \quad (5.3)$$

$$n' = (n+1) + \sum_{j=1}^q \zeta_j + q\tau \quad (5.4)$$

The value in [Equation 5.2](#) is upper-bounded by the probability over $\underline{\zeta}$ that $M(\underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1}) \neq M(\underline{h}(\underline{t}^{-1}, t'_1) + \underline{\zeta} + \tau \cdot \underline{1})$. Call this event B . B occurs only when there exists $k \in [q]$ such that $\sum_{j=1}^k (\underline{h}^{-1})_j = \lceil n'/2 - 1 \rceil$.

Claim 5.3. $\Pr[B] \leq e^{-(1-\alpha)^2 n}$

Since we argued above that $\mathbb{E}_{\underline{\zeta}}[v(t_1, M(\underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1})) - v(t_1, M(\underline{h}(\underline{t}^{-1}, t'_1) + \underline{\zeta} + \tau \cdot \underline{1}))] \leq \Pr_{\underline{\zeta}}[B]$, this claim implies the theorem. ■

Proof of Claim 5.3. Observe that

$$\Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, q, \tau}} [B] = \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, q, \tau}} \left[\|\underline{\zeta}\|_{\infty} \leq \tau \quad \wedge \quad \exists k, \sum_{j=1}^k (\underline{h}^{-1})_j = \lceil n'/2 - 1 \rceil \right] \quad (5.5)$$

$$= \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q} \left[\|\underline{\zeta}\|_{\infty} \leq \tau \quad \wedge \quad \exists k, \sum_{j=1}^k (\underline{h}^{-1})_j = \lceil n'/2 - 1 \rceil \right] \quad (5.6)$$

$$\leq \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q} \left[\exists k, \sum_{j=1}^k (\underline{h}^{-1})_j = \lceil n'/2 - 1 \rceil \right] \quad (5.7)$$

$$\leq \sum_{k=1}^q \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q} \left[\sum_{j=1}^k (\underline{h}^{-1})_j = \lceil n'/2 - 1 \rceil \right] \quad (5.8)$$

where Equation 5.6 holds because by definition the distribution of $\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, q, \tau}$ is exactly the same as $\mathcal{G}_{\varepsilon}^q$ under the condition $\|\underline{\zeta}\|_{\infty} \leq \tau$.

Observe that, by the definition of \underline{h}^{-1} and n' , we may rewrite Equation 5.8 to obtain:

$$\Pr_{\underline{\zeta}} [B] \leq \sum_{k=1}^q \Pr_{\underline{\zeta}} \left[\sum_{j=1}^k (\underline{h}(\underline{t}^{-1})_j + \zeta_j) = -b + \sum_{j=k+1}^q (\underline{h}(\underline{t}^{-1})_j + \zeta_j) + (q - 2k)\tau \right] \quad (5.9)$$

where $b = 1$ if n' is even and $b = 0$ if n' is odd. Let us consider $b = 0$ (the other case follows by the same argument). By construction there are n players at position 0, so it follows that $\sum_{j=1}^k (\underline{h}^{-1})_j = n + k\tau + \sum_{j=1}^k \zeta_j$. Therefore the RHS of Equation 5.9 equals

$$\Pr_{\underline{\zeta}} \left[n + k\tau + \sum_{j=1}^k \zeta_j = \sum_{j=k+1}^q \zeta_j + (q - k)\tau \right]$$

Since $\mathcal{G}_{\varepsilon}$ is symmetric therefore ζ_j is distributed identically as $-\zeta_j$, and combined with the above we may deduce:

$$\Pr_{\underline{\zeta}} [B] \leq \sum_{k=1}^q \Pr_{\underline{\zeta}} \left[\sum_{j=1}^q \zeta_j = n - (q - 2k)\tau \right] \quad (5.10)$$

We will prove the following lemma (notice the event in the following is an inequality rather than an equality):

Lemma 5.4. *For all $k \in [q]$ and sufficiently large n , it holds that*

$$\Pr_{\underline{\zeta}} \left[\sum_{j=1}^q \zeta_j \geq n - (q - 2k)\tau \right] \leq e^{-1.9(1-\alpha)^2 n}$$

Applying this lemma to Equation 5.10 we obtain $\Pr_{\underline{\zeta}} [B] \leq qe^{-1.9(1-\alpha)^2 n}$, which, for large n , is bounded by $e^{-(1-\alpha)^2 n}$. ■

We now turn to proving [Lemma 5.4](#). The key to proving both these lemmas is the characterization of the sum of q two-sided geometric random variables given in [Fact 2.3](#): $\sum_{j=1}^q \zeta_j$ is distributed identically to $Y - Y'$ where Y, Y' are independent $\mathcal{NB}_{q,\varepsilon}$ variables.

Proof of Lemma 5.4. By [Fact 2.3](#) it holds that $\sum_{j=1}^q \zeta_j$ is distributed identically to $Y - Y'$ as stated in [Fact 2.3](#). Furthermore, since both Y, Y' are non-negative, $Y - Y' \geq n - (q - 2k)\tau$ implies that $Y \geq n - (q - 2k)\tau$. Therefore it suffices to bound $\Pr[Y \geq n - (q - 2k)\tau]$. Furthermore, it suffices to consider just the case $k = 0$, which is the worst possible.

To summarize, it suffices to bound the probability $\Pr[Y \geq n - q\tau]$ where Y is a $\mathcal{NB}_{q,\varepsilon}$ variable. We apply the second point of [Fact 2.3](#), which says that this probability is equal to the probability $\Pr[Z \leq q]$ where Z is a binomial random variable with $n - q\tau + q$ trials and success probability $1 - e^{-\varepsilon} = 1 - \alpha$. We can apply the Hoeffding bound for binomial variables and the fact that $q = \sqrt{n}$ and $\tau = o(\sqrt{n})$ (which follows from our hypothesis that $\eta = 2^{-o(\sqrt{n})/\log n}$) to conclude that, for sufficiently large n :

$$\Pr[Z \leq q] \leq e^{-2((1-\alpha)(n-q\tau) - \alpha q)^2 / (n-q\tau+q)} \leq e^{-1.9(1-\alpha)^2 n} \quad (5.11)$$

■

Releasing the histogram leaks information. Recall that M' discretizes the input into q intervals, samples $\underline{\zeta} \leftarrow_{\mathcal{R}} \mathcal{H}_{\varepsilon,\tau,q}$, computes the perturbed histogram $\underline{h}' = \underline{h}(\underline{t}) + \underline{\zeta} + \tau \cdot \underline{1}$, and outputs the median point of \underline{h}' .

We consider a slight modification of this mechanism: in addition to outputting the facility location, it also outputs the perturbed histogram \underline{h}' . Call this modified mechanism \hat{M} , and notice that \hat{M} remains PTE: privacy holds because the perturbed histogram is $(2\varepsilon, \eta)$ -differentially private, while truthfulness and efficiency remain (with respect just to the game value, before taking into account the information cost) because the facility location output is the same as what M' would have output.

Theorem 5.5. *For any $\varepsilon, \eta > 0$, and suppose \hat{M} is run with $(2\varepsilon, \eta)$ -differential privacy. Then, for all inputs $\underline{t} \in Q^n$ and all $i \in [n]$, $\text{IC}_{\hat{M}}^\eta(\text{Id}, \underline{t}, i) > \varepsilon$.*

Proof of Theorem 5.5. \hat{M} outputs a histogram perturbed by $\underline{\zeta} \leftarrow_{\mathcal{R}} \mathcal{H}_{\varepsilon,\tau,q}$. Given any database \underline{t} , construct the histogram $\underline{h}(\underline{t})$. The following [Lemma 5.6](#) says that for any player i , there exists $t' \neq t_i$ and B such that

$$\varepsilon < \log \frac{\Pr[\hat{M}(\underline{t}) \in B] - \eta}{\Pr[\hat{M}(\underline{t}^{-i}, t') \in B]} < \text{IC}_{\hat{M}}^\eta(\text{Id}, \underline{t}, i)$$

■

Lemma 5.6. *Fix any q a positive integer, $\underline{h} \in \mathbb{Z}^q$, $i \in [q]$. Let e_k denote the k 'th standard basis vector. Suppose $\underline{\zeta}$ is sampled according to the distribution $\mathcal{H}_{\varepsilon,\tau,q}$, where τ satisfies $\Pr_{\underline{\zeta} \leftarrow_{\mathcal{R}} \mathcal{H}_{\varepsilon,\tau,q}}[\|\underline{\zeta}\|_\infty > \tau] = \eta$. Then for all $j \in [q]$, $j \neq i$, there exists $B \subseteq \mathbb{Z}^q$ such that:*

$$\log \frac{\Pr[\underline{h} + e_i + \underline{\zeta} \in B] - \eta}{\Pr[\underline{h} + e_j + \underline{\zeta} \in B]} > \varepsilon$$

Proof of Lemma 5.6. Let $B = \{\underline{h}' \mid h'_i > h_i\}$. Letting $\alpha = e^{-\varepsilon}$, we calculate that:

$$\Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{h} + e_i + \underline{\zeta} \in B] = \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{\zeta}_i \geq 0] \quad (5.12)$$

$$= \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q} [\|\underline{\zeta}\|_{\infty} > \tau] + \Pr_{\underline{\zeta}_i \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}} [\tau \geq \underline{\zeta}_i \geq 0] \quad (5.13)$$

$$= \eta + \left(\frac{1-\alpha}{1+\alpha}\right) \sum_{j=0}^{\tau} \alpha^j \quad (5.14)$$

$$> \eta + \left(\frac{1-\alpha}{1+\alpha}\right) e^{\varepsilon} \sum_{j=1}^{\tau} \alpha^j \quad (5.15)$$

Above, Equation 5.13 holds because by the definition of $\mathcal{H}_{\varepsilon, \tau, q}$ (see Definition 2.1), $\underline{\zeta}_i \geq 0$ can occur one of two ways: either we sampled $\underline{\zeta}' \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q$ and got $\|\underline{\zeta}'\|_{\infty} > \tau$ so we set $\underline{\zeta} = 0$, or else we sampled $\underline{\zeta}' \leftarrow_{\mathbb{R}} \mathcal{G}_{\varepsilon}^q$ and got $\underline{\zeta}'_i \geq 0$ and we set $\underline{\zeta} = \underline{\zeta}'$. Similarly, we can deduce that:

$$\begin{aligned} \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{h} + e_j + \underline{\zeta} \in B] &= \Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{\zeta}_i \geq 1] \\ &< \left(\frac{1-\alpha}{1+\alpha}\right) \sum_{j=1}^{\tau} \alpha^j \end{aligned}$$

Therefore, we may conclude that

$$\frac{\Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{h} + e_i + \underline{\zeta} \in B] - \eta}{\Pr_{\underline{\zeta} \leftarrow_{\mathbb{R}} \mathcal{H}_{\varepsilon, \tau, q}} [\underline{h} + e_j + \underline{\zeta} \in B]} > e^{\varepsilon}$$

■

Combining Theorem 5.2 and Theorem 5.5, we conclude that \hat{M} is *not* truthful when one uses the tradeoff utility of Equation 1.1 (for reasonable settings of ν_i).

Corollary 5.7 (Formal statement of Theorem 1.6). *Fix $\varepsilon > 0$ and $\eta = 2^{-o(\sqrt{n}/\log n)}$, and let \hat{M} be the $(2\varepsilon, \eta)$ -differentially private mechanism described above. Suppose that there exists a player i such that $\nu_i = \omega(e^{-(1-e^{-\varepsilon})^2 n})$. Let σ_0 be the strategy that always outputs 0. Then, there exists $\underline{t}^{-i} \in [0, 1]^{n-1}$ such that for all $t_i \in [0, 1]$, $\underline{t} = (\underline{t}^{-i}, t_i)$ satisfies $u_1^{\eta}(\mathbf{ld}, \underline{t}) < u_1^{\eta}(\sigma_0, \underline{t})$.*

Proof. Fix i such that $\nu_i = \omega(e^{-(1-e^{-\varepsilon})^2 n})$. Let \underline{t}^{-i} be the input guaranteed to exist by Theorem 5.2. Using the definition of u_i (Equation 1.1) and applying Theorem 5.2, we have that for all $t_i \in [0, 1]$:

$$\begin{aligned} u_i(\mathbf{ld}, \underline{t}) &< \mathbb{E}_{(s, \underline{h}') \leftarrow_{\mathbb{R}} \hat{M}(\underline{t}^{-i}, 0)} [v(t_i, s)] + e^{-(1-e^{-\varepsilon})^2 n} - \nu_i \varepsilon \\ &< \mathbb{E}_{(s, \underline{h}') \leftarrow_{\mathbb{R}} \hat{M}(\underline{t}^{-i}, 0)} [v(t_i, s)] - \nu_i \cdot \mathbf{IC}_{\hat{M}}^{\eta}(\sigma_0, \underline{t}, i) &= u_i(\sigma_0, \underline{t}) \end{aligned}$$

where we used the fact that $e^{-(1-e^{-\varepsilon})^2 n} - \nu_i \varepsilon < 0 = \mathbf{IC}_{\hat{M}}^{\eta}(\sigma_0, \underline{t}, i)$. ■

This proves that, assuming the hypotheses of Corollary 5.7, not only is there \underline{t}^{-i} such that player i would prefer not to tell the truth on some possible values of t_i (which we may view as “weakly” untruthful), but there is \underline{t}^{-i} such that player i would *always* prefer to lie about his input for all values of t_i (which we may view as “strongly” untruthful).

5.2 Releasing synopses

5.2.1 Synopsis generators reveal information.

Definition 1.7 (Restated). M is a (γ, ρ) -synopsis generator on n -player inputs with respect to a class \mathcal{C} if there is a real-valued function $P(s, F)$ such that, for all $\underline{t} \in Q^n$, $\Pr[\max_{F \in \mathcal{C}} |\overline{F}(\underline{t}) - P(s, F)| \leq \gamma] \geq 1 - \rho$.

Synopsis generators give a summary of the database that is accurate with respect to a specific set of count predicates. Intuitively, it makes sense that if the synopsis must be accurate for a very rich class of predicates, then it must also be that the synopsis reveals a lot of information about the database. This is what we formalize in the following lemma, by using the VC-dimension as a quantification of the “richness” of the class of predicates.

Theorem 1.8 (Restated). Fix any $\gamma \in (0, \frac{1}{5}), \rho \in (0, 1)$. Suppose M is a mechanism that is a (γ, ρ) -synopsis generator on n -player inputs with respect to \mathcal{C} , which has VC dimension d . Then for all $\underline{t} \in Q^n$, there exists $\underline{t}' \in Q^n$ and $i \in [n]$ such that $\underline{t}, \underline{t}'$ differ in at most $4\gamma n$ entries and such that $\text{IC}_M(\text{Id}, \underline{t}', i) \geq \min(\Omega(\frac{d}{\gamma n}), \Omega(1))$.

Proof. By the definition of VC dimension, there exists a shattering set of size d . To simplify notation, let us name the shattering set $[d]$. The definition of VC dimension implies that for every $X \subseteq [d]$, there exists $F_X \in \mathcal{C}$ such that $F_X(t) = 1$ if $t \in X$ and $F_X(t) = 0$ if $t \in [d] \setminus X$. F_X can behave arbitrarily outside $[d]$.

The idea is to use a combinatorial design to show that within a small radius of \underline{t} (i.e. by changing the values of at most $O(\gamma n)$ individuals), one can find $2^{\Omega(d)}$ other inputs T such that all $\underline{t}', \underline{t}'' \in T$ are far apart from each other. From this we deduce that there exists some $\underline{t}' \in T$ such that $M(\underline{t})$ outputs a synopsis close to \underline{t}' with probability $\leq 2^{-\Omega(d)}$. However, by the definition of a synopsis generator, $M(\underline{t}')$ outputs a synopsis close to \underline{t}' with high probability. Therefore, for some sequence of hybrid inputs between $\underline{t}, \underline{t}'$, the sum of the information leaked between successive pairs in this sequence must be $\Omega(d)$, and so one of the hybrids must have large information cost.

We proceed formally. Let $P(s) = (P(s, F))_{F \in \mathcal{C}}$ be the vector containing all estimates of counts. Let $\overline{\mathcal{C}}(\underline{t}) = (\overline{F}(\underline{t}))_{F \in \mathcal{C}}$. Let $B_\gamma(\underline{t}) = \{s \in S \mid \|P(s) - \overline{\mathcal{C}}(\underline{t})\|_\infty \leq \gamma\}$ be the γ -ball induced by \underline{t} in the output space of the mechanism.

Lemma 5.8. *There exists an absolute constant $K > 0$ such that for any $1/5 > \gamma > 0$, and for all $\underline{t} \in Q^n$, there exists a set $T \subseteq Q^n$ satisfying:*

1. $|T| \geq 2^{d'/K}$ where $d' = \min(d, \gamma n)$.
2. For all $\underline{t}' \in T$, there are exactly $4\gamma n$ coordinates i such that $t'_i \neq t_i$.
3. For all $\underline{t}', \underline{t}'' \in T$, it holds that $B_\gamma(\underline{t}') \cap B_\gamma(\underline{t}'') = \emptyset$.

We first assume the lemma is true to prove the theorem. Let T be a set as guaranteed by Lemma 5.8. By the first and third properties, there exists $\underline{t}'' \in T$ such that

$$\Pr[M(\underline{t}) \in B_\gamma(\underline{t}'')] \leq 2^{-d'/K} \quad (5.16)$$

Fix such a \underline{t}'' .

Let Z denote the set of $4\gamma n$ coordinates on which \underline{t} and \underline{t}'' differ. Now consider the hybrids $\underline{t}^{(0)}, \dots, \underline{t}^{(4\gamma n)}$ where $\underline{t}^{(i)}$ agrees with \underline{t}'' on the first i coordinates in Z , and agrees with \underline{t} on the last $4\gamma n - i$ coordinates in Z (and it agrees with both on the coordinates outside Z). Clearly $\underline{t}^{(0)} = \underline{t}$ and $\underline{t}^{(4\gamma n)} = \underline{t}''$.

Let $\text{wt}(i) = -\log \Pr[M(\underline{t}^{(i)}) \in B_\gamma(\underline{t}'')]$. We know that $\text{wt}(0) \geq d'/K$ by [Equation 5.16](#), and we know that $\text{wt}(4\gamma n) \leq \log \frac{1}{1-\rho} \leq O(1)$ because M is a (γ, ρ) -synopsis generator, $\underline{t}'' = \underline{t}^{(4\gamma n)}$, and we assume that ρ is constant.

Furthermore, by the definition of w and the information cost C , it holds that $\text{IC}(\text{Id}, t^{(i)}, j_i) \geq \text{wt}(i-1) - \text{wt}(i)$, where j_i on the LHS equals the i 'th element of Z and is the only coordinate that differs between $t^{(i)}, t^{(i-1)}$. Therefore, we deduce that:

$$\begin{aligned} \frac{d'}{K} - O(1) &\leq \text{wt}(0) - \text{wt}(4\gamma n) \\ &= \sum_{i=1}^{4\gamma n} \text{wt}(i-1) - \text{wt}(i) \\ &\leq \sum_{i=1}^{4\gamma n} \text{IC}(\text{Id}, t^{(i)}, j_i) \end{aligned}$$

Since IC is non-negative, this means there exists $i \in [4\gamma n]$ such that $\text{IC}(\text{Id}, t^{(i)}, j_i) \geq \frac{1}{4\gamma n}(\frac{d'}{K} - O(1)) = \Omega(\frac{d'}{\gamma n}) = \min(\Omega(\frac{d}{\gamma n}), \Omega(1))$. \blacksquare

Proof of [Lemma 5.8](#). Let $\underline{h}(\underline{t})$ be the histogram of \underline{t} . Let us assume that $h_1 \leq h_2 \leq \dots \leq h_d$ (for notational convenience and without loss of generality, since the names of the coordinates are immaterial and one could just rearrange them to satisfy this property).

Let $d'' \leq d$ be the largest integer such that $\sum_{i=1}^{d''} h_i \leq (1-4\gamma)n$. Either $d'' = d$, or else $d'' < d$ and we can deduce that:

$$n \geq \sum_{i=1}^d h_i = \sum_{i=1}^{d''+1} h_i + \sum_{i=d''+2}^d h_i \quad (5.17)$$

$$> (1-4\gamma)n + (d-d''-1) \frac{(1-4\gamma)n}{d''+1} \quad (5.18)$$

$$\Rightarrow d'' \geq (1-4\gamma)d \quad (5.19)$$

Here we used the definition of d'' and the fact that the h_i are non-decreasing, meaning that h_i for $d \geq i > d''+1$ must satisfy $h_i \geq (1-4\gamma)n/(d''+1)$.

Set $d' = \min(d'', 12\gamma n)$. We first construct U which is a combinatorial design over $[d']$. Namely, $|U| \geq 2^{\Omega(d')}$ and each pair $X, Y \in U$ have small intersection.

1. Initially $U = \emptyset$, so pick an arbitrary $X \subseteq [d']$ of size $d'/3$. Add X to U .
2. If there exists $X \subseteq [d']$ such that $|X \cap Y| < d'/6$ for all $Y \in U$, then add X to U , otherwise halt and output U .

It is clear from the construction that, for all $X, Y \in U$, it holds that $|X \cap Y| < d'/6$. We show that U is exponentially large:

$$|U| \geq e^{2/18^2 \cdot d'/3} \quad (5.20)$$

This is a consequence of the Hoeffding inequality. Suppose we have already added i elements to U . We show that if $i < e^{2/18^2 \cdot d'/3}$ then there exists another subset that can be added. We use the probabilistic method by showing that the probability that a random subset X of $[d']$ with size $d'/3$ does not satisfy the desired property is strictly smaller than 1.

$$\Pr_{X \leftarrow \mathcal{R} \binom{[d']}{d'/3}} [\exists Y \in U, |X \cap Y| \geq d'/6] < e^{2/18^2 \cdot (d'/3)} \Pr[|X \cap Y| \geq d'/6] < 1 \quad (5.21)$$

where we use the Hoeffding inequality (the version for sampling without replacement) in the final inequality (*i.e.* sampling $d'/3$ elements without replacement from among $[d']$, where the elements in Y are marked 1 and the rest are marked 0).

Let K be the smallest constant so that $|U| \geq 2^{d'/K}$. We now construct T using U . Let $X \in U$, then define $\underline{t}_X \in Q^n$ as follows. Let $X_1, \dots, X_{d'/3} \in [d']$ be the elements of X , say sorted in increasing order.

1. Initialize $i = 1$ and $j = 1$. (i will take value between $1, \dots, n$ and j between $1, \dots, d'/3$).
2. Do the following while $j \leq d'/3$:
 - (a) Take the first $12\gamma n/d'$ individuals after and including the i 'th individual whose values lie in $Q \setminus [d']$, and change their values to X_j .
 - (b) Set i to be the individual after the last individual modified in the previous step, and increment j .

Observe two facts: first, we never “run out” of individuals to modify, since our choice of $d' \leq d''$ and Equation 5.19 ensure that the number of players with value in $Q \setminus [d']$ is at least $4\gamma n$. Second, $\frac{12\gamma n}{d'} \geq 1$ so in each iteration we modify at least one player.

We prove that $T = \{\underline{t}_X \mid X \in U\}$ satisfies the properties of the lemma. First, it is clear that if $X \neq Y$ then $\underline{t}_X \neq \underline{t}_Y$, and therefore $|T| \geq 2^{d'/K}$. The second property holds because we modify $12\gamma n/d'$ individuals in each iteration, and there are $d'/3$ iterations.

To prove the third property, let $\underline{t}_X, \underline{t}_Y \in T$ be two distinct elements of T . Let $F_X \in \mathcal{C}$ be a function satisfying $F_X(x) = 1$ if $x \in X$ and $F_X(x) = 0$ if $x \in [d'] \setminus X$ (and F_X can behave arbitrarily outside $[d']$). Such F_X exists because $X \subseteq [d'] \subseteq [d]$ and $[d]$ is shattered by \mathcal{C} . Let $Z \subseteq [n]$ be the first $4\gamma n$ coordinates of \underline{t} taking value in $Q \setminus [d']$. Observe that \underline{t} and \underline{t}_X are identical on all coordinates outside of Z . It holds that:

$$|\overline{F}_X(\underline{t}_X) - \overline{F}_X(\underline{t}_Y)| = \frac{1}{n} \left| \sum_{i=1}^n (F_X((\underline{t}_X)_i) - F_X((\underline{t}_Y)_i)) \right| \quad (5.22)$$

$$= \frac{1}{n} \left| \sum_{i \in Z} F_X((\underline{t}_X)_i) - \sum_{i \in Z} F_X((\underline{t}_Y)_i) \right| \quad (5.23)$$

$$= \frac{1}{n} |Z| - \frac{12\gamma n}{d'} \cdot |X \cap Y| \quad (5.24)$$

$$> \frac{1}{n} (4\gamma n - \frac{12\gamma n}{d'} \cdot \frac{d'}{6}) \quad (5.25)$$

$$= 2\gamma \quad (5.26)$$

Suppose now for the sake of contradiction that $\exists s \in B_\gamma(\underline{t}_X) \cap B_\gamma(\underline{t}_Y)$. This means that $\|\overline{C}(\underline{t}_X) - P(s)\|_\infty \leq \gamma$ and $\|\overline{C}(\underline{t}_Y) - P(s)\|_\infty \leq \gamma$. But by the triangle inequality, this would imply that:

$$2\gamma < |\overline{F}_X(\underline{t}_X) - \overline{F}_X(\underline{t}_Y)| \leq |\overline{F}_X(\underline{t}_X) - P(s)| + |P(s) - \overline{F}_X(\underline{t}_Y)| \leq 2\gamma$$

which is a contradiction, and therefore $B_\gamma(\underline{t}_X) \cap B_\gamma(\underline{t}_Y) = \emptyset$. ■

5.2.2 Non-reactive mechanisms.

As in Section 5.1, we would like to use Theorem 1.8 to infer that if the database owner publishes a synopsis rather than just the outcome of the game (in the hopes that the synopsis may be useful for other purposes), then individuals may prefer to lie because their gain in information cost outweighs

their loss in value derived from the outcome. Intuitively this happens if by deviating, a player cannot lose too much value. To formalize this, let us define a mechanism M to be (β, γ) -non-reactive if there exists $\underline{t} \in Q^n$ such that for all \underline{t}' that differ from \underline{t} in at most γn coordinates, for every $i \in [n]$ and $t'' \in Q$, it holds that $\mathbb{E}[v(t'_i, M(\underline{t}'))] \leq \mathbb{E}[v(t'_i, M((\underline{t}')^{-i}, t''))] + \beta$. The following is an easy corollary of [Theorem 1.8](#).

Corollary 5.9. *Fix any $\gamma \in (0, \frac{1}{5}), \rho \in (0, 1)$. If M is a (γ, ρ) -synopsis generator for \mathcal{C} of VC-dimension d . Let $\nu = \min_i \nu_i$ and suppose that M is also $(o(\frac{\nu d'}{\gamma n}), 4\gamma)$ -non-reactive, where $d' = \min(d, \gamma n)$. Then there exists $\underline{t} \in Q^n, i \in [n]$ and a strategy $\sigma(t)$ that is independent of t such that $u_i(\text{Id}, \underline{t}) < u_i(\sigma, \underline{t})$.*

Proof. By the definition of insensitivity, let $\underline{t} \in Q^n$ be such that for all \underline{t}' differing from \underline{t} in at most $4\gamma n$ coordinates, for all $i \in [n], t'' \in Q$, it holds that

$$\mathbb{E}[v(t'_i, M(\underline{t}'))] \leq \mathbb{E}[v(t'_i, M((\underline{t}')^{-i}, t''))] + o(\frac{\nu d}{\gamma n})$$

By [Theorem 1.8](#), one of these \underline{t}' satisfies $\text{IC}(\text{Id}, \underline{t}', i) \geq \Omega(\frac{d}{\gamma n})$. Therefore, if we let σ be the strategy that outputs an arbitrary constant value in $x \in Q$ (and therefore $\text{IC}(\sigma, \underline{t}', i) = 0$), we may write:

$$\begin{aligned} u_i(\text{Id}, \underline{t}') &\leq \mathbb{E}[v(t'_i, M(\underline{t}'))] - \nu \cdot \text{IC}(\text{Id}, \underline{t}', i) \\ &< \mathbb{E}[v(t'_i, M((\underline{t}')^{-i}, x))] + o(\frac{\nu d}{\gamma n}) - \Omega(\frac{\nu d}{\gamma n}) \\ &< \mathbb{E}[v(t'_i, M((\underline{t}')^{-i}, x))] - \nu \cdot \text{IC}(\sigma, \underline{t}', i) \\ &= u_i(\sigma, \underline{t}') \end{aligned}$$

Therefore, M is not truthful. ■

In particular, [Corollary 5.9](#) holds even if M is differentially private as long as the VC-dimension of \mathcal{C} is large. Blum et al. [4] prove that it is possible for M to be ε -differentially private and still be a (γ, ρ) -synopsis generator for a class of predicates with VC-dimension $d = \Omega(\frac{\gamma^3 \varepsilon n}{\log |Q| \log(1/\rho)})$. If in addition $|Q| = \text{poly}(n)$ and $\nu = \Omega(1/\log |Q|)$ (which by [Remark 5.1](#) constitutes a realistic setting of parameters), and M is $(o(1/\log n), 4\gamma)$ -non-reactive, then M cannot be truthful.

In fact, mechanisms may be quite non-reactive because intuitively the influence of a single individual on the outcome should diminish rapidly as there are more players. As a concrete example of such a class of mechanisms, we show that any efficient mechanism for a 1-facility location game over an arbitrary bounded metric space must be non-reactive. Let (Q, \mathbf{d}) be a general bounded metric space, normalized so that $\max_{t, t' \in Q} \mathbf{d}(t, t') \leq 1$. The general 1-facility location game is defined similarly to LINE-1-FAC, except that the type space is Q rather than just $[0, 1]$.

Theorem 5.10. *Suppose M is a δ -efficient mechanism for the 1-facility location game over a bounded metric space (Q, \mathbf{d}) . Then for any $\gamma \in (0, \frac{1}{2})$, it holds that M is $(\frac{2\delta}{(1-2\gamma)^{n-2}}, \gamma)$ -non-reactive.*

Corollary 5.11. *Let $\gamma \in (0, \frac{1}{10}), \rho \in (0, 1)$. Suppose M is a δ -efficient mechanism for the 1-facility location game on a bounded metric space, and also M is a (γ, ρ) -synopsis generator for \mathcal{C} of VC-dimension d . Let $\nu = \min_i \nu_i$. If $\delta = o(\nu d')$ where $d' = \min(d, \gamma n)$, then there exists $\underline{t} \in Q^n, i \in [n]$ and a strategy $\sigma(t)$ that is independent of t such that $u_i(\text{Id}, \underline{t}) < u_i(\sigma, \underline{t})$.*

For example, the above theorem applies for the choice of parameters $d = \Omega(\frac{\gamma^3 \varepsilon n}{\log |Q| \log(1/\rho)})$, $\delta = n^{0.99}$, $|Q| = \text{poly}(n)$, and $\nu = \Omega(1/\log |Q|)$.

[Corollary 5.11](#) applies to a much broader setting than [Corollary 5.7](#) in terms of the games considered. However, even when applied to LINE-1-FAC, [Corollary 5.11](#) gives an incomparable result. Whereas [Corollary 5.7](#) applies to the specific mechanism ([Algorithm 4.6](#)) studied in this paper, [Theorem 5.10](#) holds for *any* efficient mechanism. On the other hand, [Corollary 5.7](#) is better quantitatively, and also applies when only a histogram of the discretization of the player types is released, which may contain less information than a synopsis. (One can reconstruct the histogram from a synopsis if \mathcal{C} is sufficiently rich, see for example [Theorem 4.1](#) of [\[12\]](#).)

Proof of [Theorem 5.10](#). Let $\underline{t} \in Q^n$ be the vector where all coordinates have value x for an arbitrary $x \in Q$. Fix $\underline{t}' \in Q^n$ different from \underline{t} in at most $m < n/2$ coordinates, *i.e.* m coordinates of \underline{t}' are not equal to x . Suppose for convenience of notation that these are the first coordinates t_1, \dots, t_m . We know that, by picking $s = x$, it is possible to achieve welfare $w(\underline{t}', x) \geq -\sum_{i=1}^m d(x, t'_i)$, and therefore by the δ -efficiency of the mechanism, it holds that

$$\delta \geq w(\underline{t}', x) - \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [w(\underline{t}', s)] \quad (5.27)$$

$$= -\sum_{i=1}^m d(x, t'_i) + \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} \left[(n-m) \cdot d(s, x) + \sum_{i=1}^m d(s, t'_i) \right] \quad (5.28)$$

$$= \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} \left[(n-2m) \cdot d(s, x) + \sum_{i=1}^m (d(s, t'_i) + d(s, x) - d(x, t'_i)) \right] \quad (5.29)$$

$$\geq \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [(n-2m) \cdot d(s, x)] \quad (\text{using triangle inequality}) \quad (5.30)$$

$$\Rightarrow \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(s, x)] \leq \frac{\delta}{n-2m} \quad (5.31)$$

For any $y \in Q$, we may write the following, using the triangle inequality and [Equation 5.31](#):

$$\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [v(y, s)] = -\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(y, s)] \quad (5.32)$$

$$\leq -\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(y, x) - d(s, x)] \quad (5.33)$$

$$= -d(y, x) + \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(s, x)] \quad (5.34)$$

$$\leq v(y, x) + \frac{\delta}{n-2m} \quad (5.35)$$

$$\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [v(y, s)] = -\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(y, s)] \quad (5.36)$$

$$\geq -\mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(y, x) + d(s, x)] \quad (5.37)$$

$$= -d(y, x) - \mathbb{E}_{s \leftarrow \text{RM}(\underline{t}')} [d(s, x)] \quad (5.38)$$

$$\geq v(y, x) - \frac{\delta}{n-2m} \quad (5.39)$$

If \underline{t}' differs from \underline{t} in at most γn coordinates, then for every i , it holds that $((\underline{t}')^{-i}, t'')$ differs from \underline{t} in at most $\gamma n + 1$ coordinates. Therefore we can apply [Equation 5.35](#) and [Equation 5.39](#) for $m = \gamma n + 1$ to obtain:

$$\begin{aligned} \mathbb{E}[v(t'_i, M(\underline{t}'))] &\leq v(t'_i, x) + \frac{\delta}{(1-2\gamma)n-2} \\ &\leq \mathbb{E}[v(t'_i, M((\underline{t}')^{-i}, t''))] + \frac{2\delta}{(1-2\gamma)n-2} \end{aligned}$$

■

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A Additional facts

Lemma 2.2 (Restated). *For all $i, j \in [q]$ and $U \subseteq \mathbb{Z}^q$, for $\underline{\zeta}'$ sampled from $\mathcal{H}_{\varepsilon, \tau, q}$ it holds that:*

$$\Pr[\underline{\zeta}' \in U] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' \in U] + \frac{2q\alpha^\tau}{1+\alpha}$$

where e_i denotes the i 'th standard basis vector.

Proof of Lemma 2.2. From [11, 18], it holds for all $U \subseteq \mathbb{Z}^q$ that

$$\Pr[\underline{\zeta} \in U] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta} \in U] \quad (\text{A.1})$$

Our situation is almost the same, except our perturbed histogram has the following distribution: sample $\underline{\zeta} \leftarrow_{\text{R}} \mathcal{G}_\varepsilon^q$ and check whether $\|\underline{\zeta}\|_\infty > \tau$. If so, set $\underline{\zeta}' = 0$, otherwise set $\underline{\zeta}' = \underline{\zeta}$.

Let $X_a = \{\underline{x} \in \mathbb{Z}^q \setminus \{0\}, \|\underline{x}\|_\infty \leq a\}$, the set of all non-zero points with infinity norm at most a . We will use the observation that, by the definition of $\underline{\zeta}'$, for all $\underline{x} \in X_\tau$, it holds that $\Pr[\underline{\zeta} = \underline{x}] = \Pr[\underline{\zeta}' = \underline{x}]$, and also $\Pr[\underline{\zeta} = 0] \leq \Pr[\underline{\zeta}' = 0]$.

Fix an arbitrary set U , and divide up U into three disjoint parts: $U_0 = U \cap \{0\}$, $U_1 = U \cap X_{\tau-1}$, and $U_2 = U \setminus U_0 \setminus U_1$. We reason about each of the three separately. Let $\alpha = e^{-\varepsilon}$.

1. By the definition of $\underline{\zeta}'$, $\Pr[\underline{\zeta}' = 0] = \Pr[\underline{\zeta} = 0] + \Pr[\|\underline{\zeta}\|_\infty > \tau]$. Therefore, Equation A.1 and the fact that $-e_i + e_j \in X_\tau$ imply that

$$\Pr[\underline{\zeta} = 0] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta} = 0] = e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' = 0]$$

Along with Equation 2.1, this implies that

$$\Pr[\underline{\zeta}' = 0] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' = 0] + \frac{2q\alpha^{\tau+1}}{1+\alpha}$$

2. Using Equation A.1, it holds that

$$\Pr[\underline{\zeta}' \in U_1] = \Pr[\underline{\zeta} \in U_1] \quad (\text{A.2})$$

$$\leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta} \in U_1] \quad (\text{A.3})$$

Since $U_1 \subseteq X_{\tau-1}$, it holds that the shifted set $U_1 - e_i + e_j$ is contained in the set $X_\tau \cup \{0\}$. Therefore it follows that

$$\Pr[e_i - e_j + \underline{\zeta} \in U_1] \leq \Pr[e_i - e_j + \underline{\zeta}' \in U_1]$$

which, combined with Equation A.3 implies

$$\Pr[\underline{\zeta}' \in U_1] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' \in U_1]$$

3. Since $\underline{\zeta}'$ takes range in $[-\tau, \tau]^q$, we have that

$$\Pr[\underline{\zeta}' \in U_2] = \Pr[\|\underline{\zeta}\|_\infty = \tau] \leq \frac{2q(1-\alpha)\alpha^\tau}{1+\alpha} \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' \in U_2] + \frac{2q(1-\alpha)\alpha^\tau}{1+\alpha}$$

Combining all three sets gives us

$$\Pr[\underline{\zeta}' \in U] \leq e^{2\varepsilon} \Pr[e_i - e_j + \underline{\zeta}' \in U] + \frac{2q(1-\alpha)\alpha^\tau + 2q\alpha^{\tau+1}}{1+\alpha}$$

which in turn implies the lemma. ■

B Omitted Proofs About PTE Mechanisms

B.1 Proof of Theorem 4.3

Theorem 4.3 (Restated). *There is a truthful and perfectly efficient mechanism for D-L1F $_{\gamma}$.*

Proof of Theorem 4.3. The mechanism is listed in Algorithm 4.2.

Truthfulness. Since the the players are symmetric, it suffices just to consider the truthfulness of player 1. Fix $\underline{t}^{-1} = (t_2, \dots, t_n)$. Let $\underline{h} = \underline{h}(\underline{t})$.

For all j it holds that $h_j \geq 0$. Furthermore, because player 1 is in column t_1 , it holds that $h_{t_1} \geq 1$.

The mechanism's output is the minimal $s \geq 1$ such that $\sum_{j=1}^s h_j \geq n/2$. Let s be the output of the mechanism, and we consider what happens when t_1 declares some other value t^* . Let \underline{h}^* be the histogram that is identical to \underline{h} everywhere, except $h_{t_1}^* = h_{t_1} - 1 \geq 0$ and $h_{t^*}^* = h_{t^*} + 1$. Let s^* be the minimal s such that $\sum_{j=1}^{s^*} h_j^* \geq n/2$, namely the output of the mechanism on input $(\underline{t}^{-1}, t^*)$. We analyze the following cases, using the fact that both $\underline{h}, \underline{h}^*$ are non-negative:

1. $t_1 < s$. Because for all $s' < s$ it holds that $\sum_{j=1}^{s'} h_j^* \leq \sum_{j=1}^{s-1} h_j < n/2$, it follows that $s^* \geq s$. Since $t_1 < s$, this implies that $v(t_1, s^*) \leq v(t_1, s)$.
2. $t_1 = s$: in this case player 1's utility is 0, which cannot be improved (since for this game the utility is a non-positive number).
3. $t_1 > s$: Because it holds that $\sum_{j=1}^{s^*} h_j^* \geq \sum_{j=1}^s h_j \geq n/2$, therefore $s^* \leq s$. Since $t_1 > s$, this implies that $v(t_1, s^*) \leq v(t_1, s)$.

Therefore, regardless of the value of t_1 , it holds that $v(t_1, s^*) \leq v(t_1, s)$, and therefore player 1 has no incentive to misreport his type.

Efficiency. Let s be the output of the mechanism. We prove the utility is greater for s than for all other s' .

Claim B.1. *For any histogram \underline{h} and $s = M(\underline{h})$, and for all $s' \in [q]$, it holds that $\sum_{j=1}^q h_j |j - s'| \geq \sum_{j=1}^q h_j |j - s|$.*

Recalling that $v(t, s) = -\gamma|t - s|$, this immediately implies that $w(t, s') \leq w(t, s)$ for all s' .

We now prove the claim. First consider $s' < s$. Split the summation $\sum_{j=1}^q h_j |j - s'|$ into three parts:

$$\sum_{j=1}^q h_j |j - s'| = \sum_{j \leq s'} h_j (s' - j) + \sum_{s' < j < s} h_j (j - s') + \sum_{s \leq j \leq q} h_j (j - s') \quad (\text{B.1})$$

We will bound each of the three terms. Suppose first that $s' < s$, then the following hold:

$$\sum_{j \leq s'} h_j (s' - j) = \sum_{j \leq s'} h_j (s - j) - (s - s') \sum_{j \leq s'} h_j \quad (\text{B.2})$$

$$\sum_{s \leq j} h_j (j - s') = \sum_{s \leq j} h_j (j - s) - (s' - s) \sum_{s \leq j} h_j \quad (\text{B.3})$$

$$\begin{aligned} \sum_{s' < j < s} h_j (j - s') &= \sum_{s' < j < s} h_j (s - j) - \sum_{s' < j < s} h_j (s + s' - 2j) \\ &\geq \sum_{s' < j < s} h_j (s - j) - (s - s') \sum_{s' < j < s} h_j \end{aligned} \quad (\text{B.4})$$

Input: histogram \underline{h} with m coordinates. Let $n = \sum_{i=1}^m h_i$.

1. Select a random permutation $\pi : [n] \rightarrow [n]$.
2. For $i = 1$ to n , set $t_{\sigma(i)} = \min_{j \in [m]} \sum_{k=1}^j h_k \geq i$. By abuse of notation, let $\pi(\underline{h})$ denote this setting of t_1, \dots, t_n .
3. Run $M(t_1, \dots, t_n)$.

Algorithm C.1. Converting an arbitrary mechanism to one looking only at histogram.

The first two equalities follow by definition, while [Equation B.4](#) is justified by the inequality $s + s' - 2j < s - s'$ and the fact that the h_j are non-negative.

Applying [Equation B.2](#) [Equation B.3](#), [Equation B.4](#) to [Equation B.1](#), we obtain

$$\sum_{j=1}^q h_j |j - s'| \geq (s - s') \left(\sum_{s \leq j} h_j - \sum_{j < s} h_j \right) + \sum_{j=1}^q h_j |j - s| \quad (\text{B.5})$$

Since $s - s' > 0$, $\sum_{j=1}^q h_j = n$, and, by the definition of s it holds that $\sum_{j < s} h_j < n/2$, it follows that the first term on the RHS of [Equation B.5](#) is positive. This implies the claim for the case $s' \leq s$.

The case for $s' \geq s$ follows similarly, resulting in the inequality

$$\sum_{j=1}^q h_j |j - s'| \geq (s' - s) \left(\sum_{s < j} h_j - \sum_{j \leq s} h_j \right) + \sum_{j=1}^q h_j |j - s|$$

Now, using the fact that $s' \geq s$ and by the definition of s , it holds that $\sum_{s < j} h_j \leq n/2$, we can similarly conclude that the claim also holds in this case. ■

C Mechanisms need only consider the histogram

Suppose M is an arbitrary mechanism, possibly looking at individual types. We transform it into a mechanism M' that considers only the histogram according to [Algorithm C.1](#).

M' is efficient because for all π it holds that the histogram $\pi(\underline{h})$ is exactly \underline{h} . Since the outcome of M is efficient on $\pi(\underline{h})$, therefore the outcome is efficient for \underline{h} .

To see that M' is truthful, suppose not and that there is an input \underline{h} and a deviation that allows one player to improve his utility on this input. By an averaging argument, this means that there exists π such that there is a deviation that allows one player to improve his utility for the input $\pi(\underline{h})$ and with the mechanism M . But this contradicts the truthfulness of M .